

***Projections of National Health Expenditures: Methodology and Model Specification***

The Office of the Actuary (OACT) in the Centers for Medicare & Medicaid Services (CMS) annually produces 10-year projections of health care spending for categories within the National Health Accounts (NHA). The NHA track health spending by source of funds (for example, private, Medicare, Medicaid) and by type of service (hospital, physician, pharmaceuticals, etc.). For the first time, these projections include the impact of the Medicare Prescription Drug, Improvement, and Modernization Act (or MMA), which was signed into law in December 2003.

To produce projections for total National Health Expenditures (NHE), OACT combines projections for Medicare and Medicaid spending (based on actuarial techniques) with projections for private health spending (based on a multi-equation structural econometric model, hereafter referred to as the NHE Projection Model). The NHE Projection Model attempts to capture the causal relationships between major macroeconomic variables and private health spending, as well as interactions among major causal variables within the health sector. The macroeconomic and demographic outlook from the 2004 Trustees Report and the projections of Medicare and Medicaid spending produced by OACT are exogenous inputs into the model.

Forecasting is contingent upon assumptions about macroeconomic conditions and their relationship to health care spending; thus, our projections are always subject to considerable uncertainty. As we have no historical experience with Medicare Part D, the uncertainty associated with this set of projections is greater than in previous years.

The methodology and specification for the NHE Projection Model are presented below. The discussion is organized in the following sections:

- I. Data Sources**
- II. Model Specification**
- III. Types of Services**
- IV. Sources of Funding**

## I. Data Sources

### *Health Expenditures*

All historical data for health expenditures are derived from the NHA compiled by OACT. The NHA is a national level matrix of health spending data by type of service and source of funding. Information on the methodology used in producing these historical estimates can be found at <http://www.cms.hhs.gov/statistics/nhe/definitions-sources-methods/>. Types of services and sources of funding projected in our model are listed below.

### Types of Services

#### National Health Expenditures

##### Health Services and Supplies

##### Personal Health Care

##### Hospital Care

##### Professional Services

##### Physician and Clinical Services

##### Other Professional Services

##### Dental Services

##### Other Personal Health Care

##### Nursing Home and Home Health

##### Nursing Home Care

##### Home Health

##### Retail Outlet Sales of Medical Products

##### Prescription Drugs

##### Durable Medical Equipment

##### Nondurable Medical Products

##### Government Administration and Net Cost of Private Health Insurance

##### Government Public Health Activities

##### Investment

##### Construction

##### Research

### Sources of Funding

#### National Health Expenditures

##### Private

##### Private Health Insurance

##### Consumer Out-of-Pocket

##### Other Private

##### Public

##### Federal

##### Medicare

##### Medicaid

##### Other Federal

##### State and Local

##### Medicaid

##### Other State and Local

### *Medical Price Indexes*

Data sources for medical prices are consistent with those used in the NHA. For most types of services, price indexes are based on the Consumer Price Indexes (CPI) published by the Bureau of Labor Statistics (BLS). However, for nursing home services we use input price indexes compiled by CMS, and for hospital services we use a Producer Price Index (PPI) from BLS. For inpatient hospital services in the period from 1993 forward, the NHA uses the PPI for hospital services introduced in December 1992. To obtain a measure closer to a transaction price, the PPI uses a methodology that attempts to capture discounts and redefines the “items” included in the index. For years prior to 1993, OACT estimated a transaction price measure based on an adjusted version of the CPI for hospital and related services.

For skilled nursing services, for which no separate price index is available for the time period required, we use input price indexes (IPIs) developed by OACT to track input costs incurred by these providers. IPIs are used as a proxy for output prices based on the assumption that input costs will be a major determinant of output prices. Use of the IPI implies that we fail to capture the effects on output prices of productivity change and fluctuating profit margins.

Our price measure for total personal health care spending is a chain-weighted deflator based on the indexes in the table below, with the weight for each index set equal to the share of personal health care expenditures accounted for by that type of service.

### *Derivation of the personal health care expenditure chain-type annual-weighted price index*

<b>Industry/Commodity or Service</b>	<b>Price proxy</b>	<b>2003 weight</b>
<b>Personal health care</b>		100.0
<b>Hospital care</b>	PPI, hospitals*	35.8
<b>Physician and clinical services</b>	CPI, physician services	25.7
<b>Other professional services</b>	CPI, professional services	3.4
<b>Dental services</b>	CPI, dental services	5.2
<b>Home health care</b>	CPI, professional services	2.8
<b>Other personal health care</b>	CPI, medical care	3.4
<b>Nursing home care</b>	National Nursing Home Input Price Index	7.7
<b>Prescription drugs</b>	CPI, prescription drugs and medical supplies	12.4
<b>Other non-durable medical products</b>	CPI, internal & respiratory over-the-counter drugs	2.3
<b>Durable medical equipment</b>	CPI, eyeglasses and eye care	1.4

\*Producer Price Index for hospitals, U.S. Department of Labor, Bureau of Labor Statistics. Used beginning in 1994 and scaled to 100.0 in 2000. Indexes for 1960-93 are based on a CMS developed output or transaction price index.

### *Insurance Coverage Data*

Private health insurance enrollment data are compiled by OACT using a combination of the National Health Interview Survey (NHIS) and the Current Population Survey (CPS). Presently, the insured population is benchmarked to the 1997 NHIS and is then escalated using the change in the insured population from the CPS. Net enrollment for earlier years was developed using data on gross and net enrollment and percentage of population

privately insured, as presented by M. Carroll and R. Arnett in "Private Health Insurance Plans in 1978 and 1979: A Review of Coverage, Enrollment and Financial Experience" and in preceding articles (1970-79).

Total enrollment in Health Maintenance Organizations (HMOs) is based on data from the Group Health Association of America, which later became the American Association of Health Plans (AAHP) through 1994. Data for Medicare and Medicaid HMO enrollment are compiled by OACT from CMS program data. Private enrollment is estimated as a residual after subtracting Medicare and Medicaid enrollment from total enrollment.

### *Exogenous Projections*

Projections for macroeconomic variables, such as economic growth and economy-wide inflation, and demographic variables, such as the age composition of the population, are derived from the annual projections of the Board of Trustees for OASDI (Federal Old-Age and Survivors Insurance and Disability Insurance). These projections are produced annually by the Social Security Administration (SSA).<sup>1</sup>

A projection for disposable personal income (DPI) consistent with the economic assumptions from the 2004 Medicare Trustees Report is generated using the University of Maryland Long Term Interindustry Forecasting Tool (LIFT). The relationship between DPI and GDP is influenced by fluctuations in taxes and government transfer payments, depreciation of capital stock, and retained earnings and transfer payments of private business.

The Board of Trustees for Medicare reports annually to the Congress on the actuarial status of the Hospital Insurance and Supplementary Medical Insurance Trust Funds.<sup>2</sup> These projections, as well as the Medicaid and SCHIP projections, are produced by OACT and are also consistent with macroeconomic and demographic assumptions included in the OASDI Trustees Report.

Projections for sectoral input price indexes are based on projections from Global Insight Inc. Since these projections are generated conditional on macroeconomic assumptions for aggregate wage and price growth that differ from those incorporated in the OASDI Trustees report, price and wage proxies proxy included in these indexes are adjusted for consistency with OASDI. This adjustment is based on an econometric model for the historical relationship of each index proxy to the CPI-all items and total wage compensation per employee.

The latest release of the NHE projections was produced in the fall of 2004. This forecast incorporates projections from the 2004 Trustees Reports issued in the spring of 2004, updated to reflect additional macroeconomic, Medicare and Medicaid data available through December 2004.

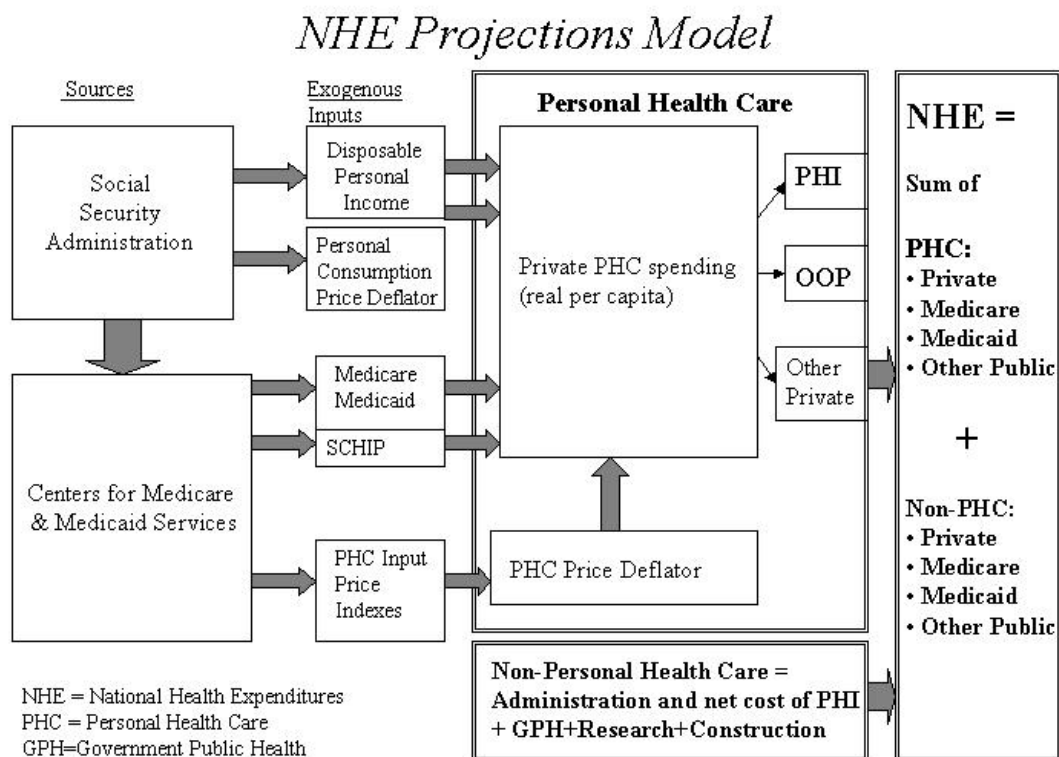
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<sup>1</sup> Board of Trustees, Federal Old-Age and Survivors Insurance and Disability Trust Funds, *2004 Annual Report of the Board of Trustees of the Federal Old-Age and Survivors Insurance and Disability Insurance Trust Funds*, <[www.ssa.gov/OACT/TR/TR04/index.html](http://www.ssa.gov/OACT/TR/TR04/index.html)>.

<sup>2</sup> Board of Trustees, *2004 Annual Report of the Boards of Trustees of the Federal Hospital Insurance Trust and Federal Supplementary Medical Insurance Trust Funds*, 23 March 2004, <[cms.hhs.gov/publications/trusteesreport/2004/tr.pdf](http://cms.hhs.gov/publications/trusteesreport/2004/tr.pdf)>.

## II. Model Specification

The structure of the NHE Projection Model for private health spending draws on standard economic theory and the health economics literature. We used these resources not only to develop the model structure, but also to evaluate the parameters resulting from our estimation. This model structure has remained largely consistent since its introduction in our projections released in July 1998. The diagram below provides a schematic view of the aggregate health sector within the NHE Projections Model and shows the linkages among the data sources, exogenous data, the personal health care (PHC) model, the non-PHC output, and the aggregate NHE projections.



The NHE Projection Model is a “top-down” model in that the growth in private health care spending and medical inflation are primarily determined at the aggregate level on the basis of exogenous projections of macroeconomic variables, actuarial projections of spending for the Medicare and Medicaid programs, and health sector assumptions. Models for spending growth and price inflation for individual types of medical services are estimated and solved separately, based on models similar in specification to the aggregate model. Model solutions for all types of services are then constrained for consistency with the aggregate spending projection.<sup>3</sup> Our choice of projection models reflects our finding that the model is substantially more robust at the aggregate level.<sup>4</sup>

<sup>3</sup> See discussion of sectoral constraints under ‘Types of Service’

<sup>4</sup> There are several possible reasons for this finding. First, spending for the different types of services is interdependent, and conceptual and measurement issues with the data make it difficult to convincingly capture complementary and substitutive relationships across types of services. When shifts across services are believed to have occurred on a large scale, it is difficult to accurately capture the effect on patterns of growth. For example, such a shift occurred following the introduction of Medicare’s prospective payment system for most inpatient hospital services. However, the manner in which such events are specified affects the coefficients obtained on the

A simple econometric model for projections of private sources of funds (private health insurance, out-of-pocket spending and other private spending) explains change in share of each source of private funding within each type of service (e.g. PHI spending on prescription drugs as a share of total spending on prescription drugs).

The core of our aggregate model of PHC spending consists of two behavioral equations:<sup>5</sup>

- Private personal health care spending (real per capita)
- Personal health care price inflation

The key variable in our aggregate model of personal health care spending is real per capita private PHC spending. Our equation for this dependent variable includes three independent variables:

- **Disposable personal income growth (less Medicare and Medicaid, real per capita)**  
(Exogenous)
- **Relative medical price inflation (PHC)**  
(Endogenous)
- **Public spending growth (PHC, real per capita)**  
(Exogenous)

The interpretation and model specification for each of these variables are discussed below.

### **Disposable Personal Income**

Income is defined as real per capita disposable personal income (DPI) less Medicaid and Medicare payments.<sup>6</sup> This is a highly influential variable in our model of private health spending. The importance of this variable is consistent with a large body of literature examining the empirical relationship between national income and health spending. It has been repeatedly shown that variations in GDP, and thus income, account for the majority of international variation in health spending. A number of studies based on time-series cross-country data for Organization of Economic Cooperation and Development (OECD) economies have largely confirmed the importance of this relationship.<sup>7</sup>

In the NHE Projections Model, income has a lagged effect on health spending. This effect is suggested by several characteristics of the market for health services. The critical element is the role of third-party payers. Since private insurers or public payers account for the large majority of health spending, spending is largely insulated from contemporaneous changes in household income. Furthermore, since consumers generally do not pay for their medical expenses directly at the point of purchase, their decisions are not immediately affected in the short term by variations in income.<sup>8</sup>

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model variables. Second, data on relative prices across types of medical services are somewhat flawed for our purposes and are not always consistent across services; thus, obtaining reasonable cross-price elasticities is difficult. Third, health services tend to be purchased as bundles that incorporate types of services extending across several different sectors, while the data are not measured in such a way that we can track the behavior of the market for these linked bundles. Aggregation across all types of medical care ameliorates these problems.

<sup>5</sup> Variables are expressed as log differences (growth rates)

<sup>6</sup> The objective is to obtain as nearly as possible a measure of income that applies to the population that accounts for private spending on medical care. Thus we exclude spending for Medicare and Medicaid, which are included in DPI but accrue to a population that is primarily publicly insured.

<sup>7</sup> For a review of this literature, see Gerdtham, Ulf, "International Comparisons of Health Expenditure," *Handbook of Health Economics* (Amsterdam, Elsevier, 2000): 11-53.

<sup>8</sup> Some current period effect can be expected in response to consumer cost-sharing and loss of employment, with the

Conceptually, the effect of income on private health spending could be affected by the decision to purchase private health insurance. Increases in income would encourage the purchase of more generous insurance. However, the current-period response is also dampened by the intermediation of employers. Most insurance is purchased through employers, who respond to the pooled interests of their employees. Employers often offer few choices of health plans, and some offer none, limiting short-term flexibility for employees. The introduction of new insurance options in response to employee preferences can be expected to occur with a lag. Negotiating health insurance contracts on an annual basis also causes a delay in the time it takes to respond to changing labor market conditions and employee preferences. In addition, because the exercise of control over medical expenses by private insurers may require the development of new organizational structures (for example, managed care organizations or health savings accounts), such changes may require alternation in government regulation of institutions or in the legal environment. Therefore, several years may be required before a response to changes in income can be fully realized.

Public spending decisions can also be expected to respond to changes in income with a substantial lag. Changes in the growth of public spending will be influenced by the underlying health sector variables that drive the cost of services and by changes to the regulations that affect the price and volume of these services. Examples include the incentive effects of the physician fee schedule or prospective payment systems, Federal and state level regulation influencing the nature of insurance coverage (e.g. diverse forms of “patient protection” legislation), or costs associated with medical malpractice liability. Such changes occur over time as lawmakers respond to perceived problems in the financial status of the programs within the limits of what taxpayers are willing to pay for them.

To capture these potential lags, the income term in our model of personal health care spending is incorporated as a moving average over 5 years (from four years previous through the current period).<sup>9</sup> The sum of the coefficients for all lagged periods on this variable is 1.95, which, given our specification, can be interpreted as an income elasticity. This elasticity implies that a 1-percent increase in real disposable personal income, less Medicare and Medicaid spending, results in a 1.95-percent increase in private personal health care spending spread over a period of five subsequent years. This compares with macro-level elasticities of approximately 1.0 to 1.5 in the empirical literature.<sup>10</sup> However, these estimates are generally based on spending by all sources of funding, rather than on private spending alone, and are estimated based on international cross-country time-series data sets. Given the absence of an explicit measure of technological change, it is also likely that this coefficient on income captures an interaction effect between income growth and medical innovation.<sup>11</sup>

As discussed above, this income term is intended primarily as a proxy for the influence of a number of separate developments, which occur in response to changes in incomes. These might include changes in the nature and breadth of health insurance coverage offered by employers, the development and evolution of institutional structures for the financing and delivery of medical care (e.g. including the legal environment and organizational structures within the private sector that facilitate the development and diffusion of new forms of coverage), shifts across different forms of managed care, the passage of state and Federal legislation influencing the costs of providing care, and fluctuations in the fraction of the population with health insurance.<sup>12</sup> The implicit theory

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associated loss of employer-provided health insurance.

<sup>9</sup> Estimates that allow coefficients to vary across this five-year period based on a polynomial distributed lag (PDL) show no statistically significant improvement in explanatory power over a moving average.

<sup>10</sup> Getzen, T.E., “Health Care is an Individual Necessity and a National Luxury: Applying Multilevel Decision Models to the Analysis of Health Care Expenditures,” *Journal of Health Economics*, 2, 2000: 259-270.

<sup>11</sup> A constant term (sometimes interpreted as a deterministic trend associated with technological change) was found in our model to be negative and insignificant and was therefore dropped.

<sup>12</sup> Explicit measures of managed care based on the inclusion of private sector HMO coverage rates as a proxy for the effects of managed care were found to be insignificant in our model. We believe that this result was associated with problems in the breadth and consistency of the proxy, which does not capture the effects of shifts across forms

underlying this variable is that the income effect occur indirectly in the form of changes to the institutions within which medical care is provided, rather than at the level of the individual consumer, and that the specific nature of this change cannot be predicted and, almost always, cannot be accurately measured.

The projection of the income variable is exogenous (an outside input) to our model. Projections of real per capita disposable personal income incorporated in our projections are consistent with exogenous OASDI projections for growth in GDP and the economy-wide personal consumption deflator (see Data Sources section for modeling of link between GDP and DPI). Projections for Medicaid and Medicare spending (subtracted from disposable personal income) are also exogenous, based on OACT projections (see the Data Sources section for a description of all exogenous inputs and source citations).

### **Relative Medical Price Inflation**

Economic theory predicts that consumers allocate their spending on goods and services according to their relative price. However, the existence of third-party payers for medical care again complicates this relationship. Consumers bear only a fraction of the actual price of medical services at the time of purchase. Thus, in short-term consumption decisions, they respond to the out-of-pocket price rather than to the actual price, generally determined by a combination of deductibles, cost-sharing requirements, and out-of-pocket maximums. The price to consumers can be roughly approximated by the fraction of total costs paid out-of-pocket multiplied by the actual price. This was the price variable originally included in our model. We would expect this variable to appear in the model with a negative coefficient.

Medical prices also influence demand for services in two additional ways. First, the price of health insurance is effectively the price of the bundle of medical goods and services an enrollee is expected to consume (plus administrative costs and profits). Thus, consumers' decision to purchase health insurance, through their employers as agents in most cases, and the amount of health insurance purchased, is influenced by the relative price of medical care through its effect on the price of insurance. Demand for health care, therefore, depends upon changes in the actual relative price of medical care as well as the relative price on an out-of-pocket basis. Second, the relative price of medical care affects demand for services across types of medical care through the price sensitivity of health insurers' coverage and provider selection decisions. The nature of this relationship also suggests the possibility of a lag in the response to relative price.

We found that the coefficient on relative price to consumer on an out-of-pocket basis was negative (as expected) but statistically insignificant in our aggregate model. The primary reason for this finding is probably that we are working with data at a high level of aggregation. The out-of-pocket price to consumer varies substantially across different consumers, which masks the predicted effect in the aggregated data. Another reason is that the price data are somewhat flawed for our purposes—excluding discounts in a few cases—and failing to adjust completely for changes in quality, thus obscuring the relationship between quantity and price.

The variable included in our model is relative medical price inflation unadjusted for changes in consumer out-of-pocket share. This variable was found to have the predicted negative relationship with real per capita private PHC spending, and was significant at the 1-percent level. The estimated relative price elasticity is  $-0.4$ , suggesting that a 1-percent increase in the relative price of medical care results in a 0.4-percent decline in real per capita PHC spending. This elasticity is above micro-based estimates of price elasticity of demand for medical care ( $-0.1$  to  $-0.2$  from the Rand Health Insurance Experiment).<sup>13</sup> This discrepancy reflects the fact that micro-based studies use household-level data on the relationship between consumer out-of-pocket spending below out-of-pocket maximums

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of managed care over time, changes in the nature of managed care within models of managed care, or spillover effects.

<sup>13</sup> Manning, W.G., et al., "Health Insurance and the Demand for Medical Care: Evidence from a Randomized Experiment," *American Economic Review*, Vol. 77, No. 3, June 1987.



and effective price given coinsurance rates. In addition to issues associated with the use of aggregated data, such estimates do not include sensitivity to changes in insurance premiums (which reflect medical price inflation for all care delivered).

In the absence of a variable to control explicitly for declines in out-of-pocket share, we have attempted to adjust for the predicted effects of changes in this variable through judgmental adjustments to the model solution.

Medical price inflation is an endogenous variable in our model since it is determined within the NHE Projection Model. In the long term, the growth in the price of medical services and goods should be determined by growth in the relative price of inputs used in production of health services, economy-wide price inflation, changes in profit margins, and the rate of change in total factor health sector productivity. Only the first of these is explicitly included in our model.

The dependent variable in our model is OACT's price deflator for personal health care spending. This is a function of input price inflation<sup>14</sup> (IPI) over the past three years. Coefficients for lagged IPI are fitted along a second-order polynomial, with more than 80 percent of the effect estimated to occur within a year. The effects of other factors (economy-wide price inflation, productivity growth, industry profitability) are captured indirectly through their influence on IPI, and through a first-order autocorrelation adjustment.

Given the potential for price discounting associated with managed care to influence medical price inflation, we also attempted to include an HMO proxy variable in our model for this inflation. However, we found that this variable did not have a statistically significant effect – probably due to flaws in this variable as a proxy for managed care. Problems also may exist because of flaws in our price measures.

Input price inflation is an exogenous input to our model. Projections of wage and price proxies that feed into input price indexes are generated by Global Insight Inc., and are adjusted for consistency with OASDI macroeconomic assumptions.

### **Public Health Spending (PHC)**

In our model of growth in real per capita private spending on PHC, growth in real per capita public spending has a negative coefficient. Public and private sector spending are jointly affected by a number of factors. From the supply side, it is probable that the growth rates in the per enrollee cost of providing treatment to persons insured under public and private programs correspond quite closely in the long term – although demographic and institutional differences will produce some variation.

The negative coefficient on this variable in our model reflects in part that neither public nor private spending is expressed in per enrollee terms. Rather, spending is on a per capita basis – the denominator is total population. The reason for this choice lies in data issues with time series on insured population (private, Medicare, and Medicaid). The time series for private enrollment is defined to include all persons with private coverage. This including Medigap policies, where the primary source of coverage is Medicare. Thus, there is a substantial overlap between the series. In addition, the history for private enrollment stems from multiple sources and is subject to inconsistencies over time due to variations in survey questions. A second issue is that the history of Medicaid enrollment is volatile, due to changes in eligibility for the program. These changes tend to involve the influx of a relatively low per enrollee population (e.g. children and pregnant women) relative to the existing Medicaid population (which is relatively heavily weighted towards the institutionalized). This distorts per enrollee growth. The use of growth on a per capita basis measures means that a shift in enrollment between public and private programs will be associated with a change in per capita spending in the same direction – implying a negative

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<sup>14</sup> The input price index used is a weighted average of OACT's input price indexes for hospital services, physician services, home health services, nursing home services, and pharmaceuticals.

coefficient on public spending in our model.

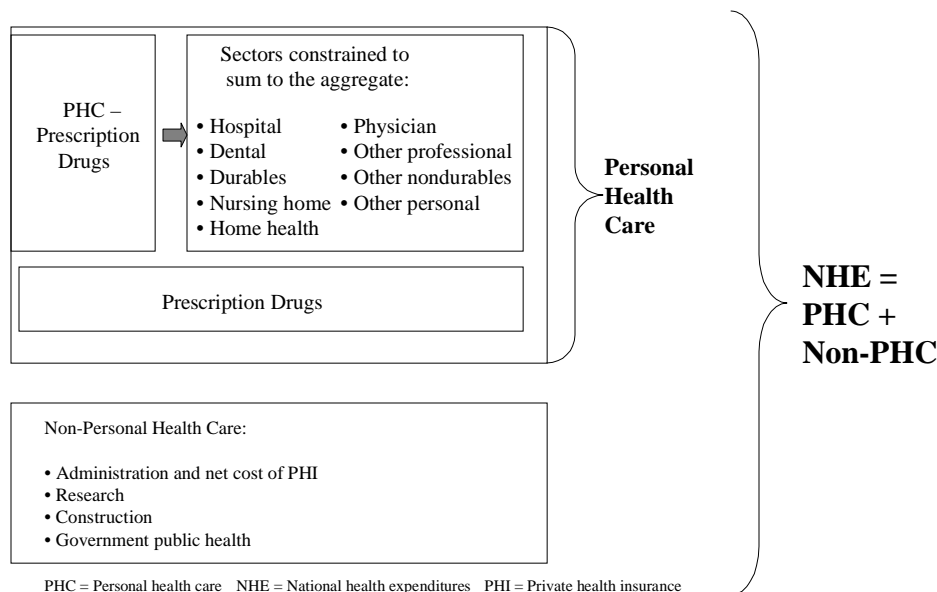
Our model forecasts private spending growth conditional on Medicare and Medicaid spending projections based on the 2004 Trustees Report. We incorporate real per capita public spending as an independent variable in our model of private spending. As discussed above, the balance of the effect of changes in public medical spending on private spending, after accounting for the shared influence of income, medical price inflation, can be expected to be negative. The negative coefficient on public spending in our regression captures two combined effects over the historical period: (1) shifts in coverage out of private and into public programs, and (2) any potential short-term cost shifting between public and private programs.

Public health spending is largely exogenous to our model – based on actuarial projections of Medicare and Medicaid spending, which account for approximately 75 percent of public spending. Medicaid and Medicare spending projections are based on OACT projections. Other public spending is projected endogenously, largely based on current and lagged growth in GDP from OASDI projections (see the Data Sources section for a description of all exogenous inputs to our model and source citations).

### III. Types of Services

Models for individual sectors of the NHE Projections Model are discussed below. Sectors are broken into personal health care (PHC) and non-personal health care (Non-PHC) categories.

#### Sectoral Composition of NHE Projections Model



The aggregate model for personal healthcare (PHC) described in the previous section is composed of ten types of services. Nine of these sectors are constrained so that the sum of types of services is equal to the model solution for total PHC. The exception is the prescription drug sector, which is not constrained because of the need to adjust our projections based on outside research not explicitly incorporated in the model. In general, the aggregation of the unconstrained model solution for the sectoral equations results in a total which is close to the aggregate model projection, so that the normalization process does not involve a major adjustment to the patterns of growth for the individual types of services. The non-personal health care components are unconstrained, but account for a much lower share of NHE than the personal health care components.

Within the constraints imposed by the aggregate model forecast of private spending and medical inflation, we use sector-specific equations to establish projections for real per capita spending growth and relative price inflation for each type of service that make up private health care spending. Only minor changes in specification were made to individual sectoral models, which remain largely unchanged from previous versions.

#### Real per Capita Spending Growth

Models of real per capita spending growth for the types of services that make up personal health care follow the specification for the aggregate model of PHC. For most sectors, these models have a specification similar to that used for aggregate personal health care. Key structural variables are:

- Disposable personal income growth (less Medicare and Medicaid, real per capita)
- Relative price inflation for the sector

- Public spending growth for the sector (real per capita)

Key differences across the models for different types of services are the lag on the income term, the relative importance of the three variables, the inclusion of dummy variables to capture phenomena specific to the sector, and in a few cases, the inclusion of additional independent variables where relevant.

Exogenous inputs to these models parallel the aggregate: real per capita disposable income less Medicare and Medicaid, and sector specific projections of Medicare and Medicaid spending.

The lag on the income term generally varies with the share of spending that is accounted for by consumers' out-of-pocket expenses: the greater the out-of-pocket share, the shorter the lag, as consumers respond more quickly to changes in their income. We evaluated coefficients on income and price terms for consistency with the aggregate regression results and across sectors; however, the relationship is not precise.

The table below summarizes the independent variables used to model real per capita spending growth for each of the personal health care sectors. For the sectors with the greatest share of NHE, we have provided some additional descriptive information about their sector models.

Sector	Dependent variable	Independent variables
<b>Hospital services</b>	Real private hospital services per capita	Real disposable personal income (lagged 5 years) (+) Relative price(-) Public spending growth (-) Dummy, (1984-)
<b>Physician services</b>	Real private physician services per capita	Real disposable personal income (lagged 5 years) (+) Relative price (-) Public spending growth (-) Dummy, 1983-85 (+)
<b>Other Professional services</b>	Real private other professional services per capita	Real disposable personal income (no lag) (+) Relative price (-) Public spending growth (-) Lagged physician spending (+) Dummy, 1986-88 (+)
<b>Prescription Drugs</b>	Real aggregate drug spending per capita*	Real disposable personal income (lagged 3 years) (+) Relative price * Share paid out-of-pocket (-) New drug introductions (+) Direct-to-consumer advertising (+)
<b>Over the Counter Drugs and Other Nondurables</b>	Real private other nondurables spending per capita	Real disposable personal income (lagged 2 years) (+) Relative price (-) Lagged dependent (+)
<b>Durables</b>	Real private durables spending per capita	Real disposable personal income (lagged 2 years) (+) Relative price (-) Public spending growth (-)

<b>Dental services</b>	Real private dental services per capita	Real disposable personal income (no lag) (+) Relative price (-) Dummy, 1981 (+)
<b>Nursing home services</b>	Real private nursing home services per capita	Public spending growth (-) Demographic index (+)
<b>Home health services</b>	Real private home health services per capita	Public spending growth (+) Demographic index (+) Dummy, 1988 (+)

\*The prescription drug model needed to be on an aggregate rather than a private basis because the shifts in payments predicted due to MMA make it impossible to accurately model real private prescription drug spending per capita over the next ten years. See the Prescription Drug section below.

### Hospital Services

Real per capita growth in private hospital spending is well explained by the variables in our template model specification. Given the low out-of-pocket share on average for hospital services (inpatient and outpatient), we anticipate a long lag between a change in household income and the time of impact on hospital spending. Our results are consistent with this expectation we estimate coefficients on lagged income growth with a polynomial distributed lag, which indicates the peak effect of income fluctuations occurs with a lag of 3 to 4 years.

The estimated price elasticity of demand is above that for the PHC, significant at the 1% level, and fairly robust to changes in regression interval. Attempts to fit an out-of-pocket variable, either in combination with the price term (i.e. effective price to consumers) or separately, were unsuccessful. However, in the hospital sector, this share is low and fairly stable (just over 3 percent for 1995 through 2003) so effects are likely to be small.

Public real per capita spending has negative coefficient as expected, capturing shifts in enrollment between private and public coverage, as well as any short-term cost-shifting effects between private and public payers.

Our current model for real per capita growth in hospital spending does not include an explicit effect for managed care. Managed care is expected to continue to influence growth in hospital spending through both utilization and relative price inflation. However, we have found that our previous proxy for managed care effects to be increasingly flawed for this purpose. The combined effect of managed care and the Medicare prospective payment system (PPS) for this sector is represented in the current model as a one-time structural change beginning after the introduction of PPS (from 1984). Current projections include a positive adjustment factor to growth in real per capita hospital spending growth, reflecting diminishing potential for additional reductions in inpatient utilization and relative price inflation relative to the historical period since 1984. Our model for relative hospital price inflation continues to include private sector HMO penetration (change and level) as a proxy for managed care, but also includes adjustments to the resulting projection.

### Physician Services

The estimated lag structure for the income term in the physician model indicates an effect which extends over five years, but is evenly weighted across periods (effectively a shorter average lag as compared with the hospital model). The sum of coefficients on all lags of the income term is substantially smaller than for the hospital sector, close to the coefficient in the aggregate model for PHC. Relative price inflation fits only weakly in this model as compared with PHC, and growth in real per capita public spending on physician services has a smaller estimated negative effect than the aggregate model.

In general, our template specification fits real per capita growth in physician spending somewhat less well than hospital spending. This primarily reflects two outlying periods: much higher than predicted growth in 1984-85, and much lower than predicted growth in 1993-96. Absent these periods, the pattern of growth implied by the income and relative price term produces a fairly good fit. Through 1983, the physician share of personal health care spending remains close to flat, drifting slightly downwards (from 24.0% in 1965 to 22.0% in 1983). From 1984 through 1994, the share rises, reaching 26.0% by 1992 before beginning to move downwards to 25.4% by 2001. Without some control for the period of rapid growth in the early 1980s, it is difficult to obtain a model with acceptable fit and reasonable coefficients.

We have included a dummy variable to capture the period of rapid growth from 1983 through 1985, while the faster growth later in the decade is consistent with the income term. Our interpretation of this variable is that it captures a non-recurring substitution effect of professional services for inpatient care. This period saw a major shift in provider incentives associated with the introduction of inpatient PPS under Medicare (spillover effects for private spending) and the initial surge in managed care enrollments. The effect of the inclusion of this dummy is that the resulting model will tend to project a pattern of growth for physician services that is more consistent with the near-stable share of PHC in the pre-1984 and post-1994 data rather than the more rapid growth of the mid-1980s.

### **Prescription Drugs**

Prescription drugs differ in important ways from other types of medical care. First, it is a product, not a service, so the cost structure of the industry differs substantially from sectors such as hospital, physician, or nursing home, where labor costs play a critical role in driving price. Second, historically, prescription drug spending has had a much larger consumer out-of-pocket share than other types of medical care, so that demand tends to be more sensitive to price. Third, the public sector has historically played a relatively small role in funding prescription drug spending. We also have access to additional information on supply and demand factors for this sector, in the form of data on new drug introductions, research spending, patent expirations, and direct-to-consumer (DTC) advertising. As a result, our model for prescription drugs is somewhat different from those developed for other sectors.

Our income variable fits with a shorter lag than we include in our aggregate model. This is the expected result based on the larger share paid on an out-of-pocket basis historically. Relative price inflation times has a strong fit and we recently multiplied this variable by the out-of-pocket prescription drug share to account for the fact that consumers' out-of-pocket share has declined steadily over the last twenty years. Public spending growth is not included as a variable in this model due to its relatively minor role.

Patterns of growth over the most recent ten years of data are by far the most difficult to explain as the effects of several different factors must be disentangled. The out-of-pocket share of spending by consumers dropped sharply as privately insured patients moved into managed care plans that generally have lower co-pays (this phenomena largely did not apply to Medicare beneficiaries, who continued to pay a relatively large share of drug costs out-of-pocket). Also, changes to regulations in 1997 dropped some of the earlier restrictions on television advertising for prescription drugs. In addition to income and relative price terms, our model for real per capita drug spending includes a four-year moving average of the number of new prescription drugs introduced, and the change in DTC advertising lagged one year.

In the most recent projection, we used the latest estimates of Medicare Part D from the Fiscal Year 2006 President's Budget to help us create a drug projection that included the impact of MMA. Since the start of Part D in 2006 is expected to shift the source of payment for drugs for Medicare beneficiaries from the private sector and Medicaid to Medicare, our first step was to change our dependent variable from real private drug spending per capita to real aggregate drug spending per capita. Once we generated a 10-year projection of total drug spending, our office generated updated Part D cost estimates based on the growth rates from our aggregate drug projection. We used these updated cost estimates to forecast shifts in sources of payment when Part D begins in 2006.

In addition to forecasting shifts in funding, we had to determine if there would be an effect on total drug spending due solely to the start of the Medicare drug benefit. There were two major offsetting factors to consider: (1) price discounts, especially for those previously uninsured for drugs or in Medigap plans, and (2) drug utilization increases because more generous drug coverage induces additional spending on drugs. Based on micro model prescription drug data, we determined that the effect of induction would be slightly higher than the effect of price discounts. The result was estimated to be approximately \$1.0 billion of additional spending on prescription drugs in 2006 because of the start of the Medicare drug benefit. Since MMA authorized penalties for late enrollment and because extensive outreach efforts are planned, we do not expect that Part D will have a notable impact on drug spending growth after 2006.

### Relative Price Inflation by Type of Service

Price inflation for individual types of services was initially modeled based on relative input prices. However, we were unable to obtain theoretically consistent and significant results, probably due to differences across types of services in the definition of both the output price and input price indexes.

Our current model attempts to explain variations in sectoral price inflation relative to personal health care (which is in turn driven by projections of input price inflation). Changes in public policy that could be expected to influence relative prices differentially across NHA sectors (such as the imposition of price controls in the early 1970s and the introduction of the prospective payment system for Medicare hospital inpatient services in 1983) are captured through the use of dummy variables. Our managed care proxy was also included in selected models (hospital, dental) since price discounting could have differing effects on price inflation for different medical services.

Variables included in models of relative price inflation are shown below:

Sector	Dependent variable	Independent variables
<b>Hospital services</b>	Hospital price inflation (relative to all medical services)	HMO Penetration (change and level, -) Dummy, 1981 to 1985 (+) Dummy, price controls, 1973 to 1974 (+)
<b>Physician services</b>	Physician services price inflation (relative to all medical services)	Dummy, price controls, 1972 to 1974 (-)
<b>Other Professional Services</b>	Other professional services price inflation (relative to all medical services)	Dummy, price controls, 1972 to 1974 (-) Physician price inflation (-)
<b>Drugs (Prescription)</b>	Drug price inflation (relative to economy-wide)	Relative input price inflation (lagged one year, +) Growth in drug research spending (4 years, +) Dummy, 1993 forward (-)
<b>Dental services</b>	Dental price inflation (relative to all medical services)	Change in HMO penetration (lagged one year, +) Dummy, 1976 Dummy, price controls, 1973 to 1974 (+)

		Dummy, 1981 to 1985 (-)
<b>Nursing home services*</b>	Nursing home input inflation	
<b>Home health services</b>	Home health price inflation	PHC price inflation

\* Projection for nursing home input price inflation is adjusted from projections by Global Insights Inc. for consistency with SSA projections of economy-wide inflation.

Generally, it proved more difficult to achieve a good fit for the relative price regressions for individual sectors than for the real per capita spending regressions. This reflects the combination of flaws and inconsistencies in the price data, and the difficulty in capturing the effects of government policy and institutional change on relative price across types of services. For example, managed care can be expected to influence prices in some sectors (e.g. hospital services) much more than others (dental services), with potentially important effects on relative price inflation for these sectors, but our proxy for capturing managed care effects is flawed. Dummy variables, are, of course, an imperfect tool for capturing effects of government policy which has effects on relative price (e.g. price controls).

Note, however, that where the regression fit for relative price inflation is not good (e.g. physician services) the resulting equation will generate a forecast which tends to track the price inflation forecast for the denominator, which often accounts for a very high fraction of variation. For example, for physician services, the denominator is PHC price inflation, which accounts for 88% of variation in price inflation for physician services from 1960 through 2003.



#### IV. Sources of Funding

In the current version of our model, we introduced econometric models for change in the composition of private spending across sources of funds (private health insurance, out-of-pocket, and other private). These models attempt to systematically capture some of the trends that had previously been introduced on a judgmental basis following the solution of the model.

Trends in insurance coverage (private, Medicaid, and Medicare enrollment, and the uninsured population) matter since the fraction paid out of pocket differs substantially across these groups. Shifts in enrollment can be expected to have an effect that varies across sectors – the most prominent example being prescription drugs. Changes such as the historical spread of managed care (with the associated pattern of much lower copayments) can also be helpful in explaining history. In addition, current period growth in disposable personal income may have an impact on out-of-pocket spending through its influence on discretionary medical spending.

Our model for private sources of funds is a hybrid between “top-down” (constrain sectors to total) and “bottom-up” (sum across sectors to obtain total) approaches. This reflects an evaluation of the tradeoffs involved. For example, the “top-down” approach eliminates the need to explicitly control for shifts across sectors, and reduces (to some extent) problems with noise in the data. However, a major disadvantage of the top-down approach is that the additional level of aggregation involved may obscure relevant trends, and make it more difficult to adjust projections precisely at the level where information is most available. For example, trends in OOP for prescription drugs have been very important in recent years because the out-of-pocket share of drug spending has, temporarily at least, stopped declining due to factors like three-tiered copayments. Prescription drugs, physician services, nursing home care, and dental services account for about two-thirds of out-of-pocket spending, but are driven by a different mix of factors. In addition, since OOP share differs markedly across sectors, shifts between sectors (for example, hospital to drug spending) will matter. It is easier to capture these shifts by summation than econometrically at the aggregate level.

Our model includes equations for out-of-pocket and other private spending as a share of total spending for both total PHC and for each type of medical product and services. To facilitate consistency across sectors, the projection for PHC is included as an independent variable in sectoral equations. However, the published forecast is set equal to the summation across all sectors. Private health insurance spending for each sector is set as a residual based on forecasts of private spending in the sector and shares of OOP and other private of this total.

The dependent variable in these models is the change in OOP (or other private) spending as a share of total spending.

Our model for change in OOP share for PHC is driven by the composition of health insurance enrollment (private, Medicare, and Medicaid) as a share of population.

Other variables included in the sectoral models include:

- OOP share for PHC
- HMO penetration rate
- Private health insurance enrollment as a share of the population
- Medicare and Medicaid enrollment as a share of the population
- Real disposable personal income

The projections produced by these models are then adjusted based on an evaluation of the model fit and (where available) on additional sources of information (for example, survey results with information on the nature of out-of-pocket payments for employer-provided health coverage).

In addition to our model of private sources of funds, we also project sources of public funds other than Medicare

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and Medicaid. These sources account for approximately 25 percent of total public spending. Other federal and other state and local spending (exclusive of Medicare and Medicaid spending) are projected based on econometric models similar to those used to project real per capita private spending models. Like our models for private spending, we specify aggregate PHC spending for other federal and other state and local, and then establish sector-level spending within the constraint of the aggregate projection.

Our projection process combines to give us a sound and defensible projection methodology based on accepted econometric and actuarial projection techniques. As with any projection, we are constantly reviewing the accuracy of our projections and working to make improvements in the methodology. Please e-mail [DNHS@cms.hhs.gov](mailto:DNHS@cms.hhs.gov) with any comments, feedback, or suggestions on our NHE Projection Model.